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Will 'we' continue to exercise? The associations between group identification, identity leadership, and relational identification on group exercise class adherence

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Abstract

The present study sought to provide support for the roles of both the social identity approach (group identification and identity leadership) and relational identification in adherence to group exercise classes. One hundred and twenty U.K. based group exercise class attendees completed an online survey in a cross-sectional design. Group identification, perceptions of identity leadership behaviour, and relational identification were not significantly associated with length of class attendance. However, both identity leadership behaviour of the class instructor and ingroup affect (measured as an element of multi-dimensional group identification), were significantly associated with expected likelihood of continued class participation. For every 1 unit increase in the identity leadership and ingroup affect measures respectively, participants were 1.9 and 2.0 times more likely to be 'at or near 100%' likely to continue vs not. Social identity variables were not associated with how long exercisers had been attending class, but identity leadership and ingroup affect are positive influencers of expected future adherence. Further research into the social identity approach and identity leadership behaviour regards exercise class attendance tenure vs volume is warranted and is likely to benefit from the use of multi-dimensional group identification measures.

KEYWORDS: Group Dynamics, Health, Behaviour Change, Physical Activity

29 **Will 'We' Continue to Exercise? The Associations between Group Identification,**
30 **Identity Leadership, and Relational Identification on Group Exercise Class Adherence**

31 Physical activity is associated with a wide range of health benefits, including
32 reduced risk of cardiovascular disease, type 2 diabetes, cancers, and depression [1]. While it
33 is recommended that adults engage in 150 minutes of moderate intensity exercise each
34 week, 41% of those aged 40-60 are undertaking nowhere this level, failing to walk briskly
35 for even 10 minutes a month [2]. Adherence to new exercise programs is also poor, and
36 around half of those who begin exercising drop out within the first six months [3]. The
37 physical and psychological benefits of regular exercise make encouraging participation and
38 understanding factors linked to adherence of great importance. There is a need to ensure
39 that adults not only start exercising, but also continue to participate. Group exercise classes
40 have been highlighted as a potential gateway to participation, with almost one third of
41 people attending having been previously inactive [4]. With this in mind, group exercise
42 classes offer a valuable avenue for research to advance our understanding of how continued
43 attendance in these settings might be encouraged.

44 **The Social Identity Approach**

45 It has been suggested that the social identity approach provides crucial and unique
46 social processes to understand and influence physical activity participation [5-7]. While the
47 social identity approach has been applied across a variety of contexts including within
48 organisations [8], and sport [5,9], research with regard to its relevance to exercise remains
49 limited (exceptions [6,7]). The social identity approach comprises two theories; social
50 identity theory [10] and self-categorisation theory [11]. Social identity theory posits that
51 individuals have personal identities and also social identities related to the groups (e.g., an
52 exercise class) to which they see themselves as part. A social identity may be considered as a
53 person's sense of self that is based on their group membership(s). Self-categorisation theory

54 describes the *process* by which individuals' move from a personal to a social identity ('me' to
55 'we') termed 'depersonalisation'. Once 'depersonalised' it is suggested that an individual
56 will reference the values and norms of the group (as opposed to their own) to inform their
57 behaviour. In this way, it is possible to hypothesise that by identifying as part of a group that
58 values regular exercise (e.g., an exercise class) an individual's behaviour may be affected in
59 the form of increased adherence. Accordingly, group (social) identification, or an
60 individual's sense of belonging or commitment to a group, may relate to length of class
61 participation and/or intention to continue to participate. This is not of course to suggest that
62 personal identity is unimportant in this regard. Indeed, running role identity may be more
63 important than running group identity in terms of adherence in the face of group disbandment
64 [12], and the development of group and personal identities in relation to exercise may be
65 intertwined [13]. Examination of this however, falls beyond the scope of the present study.

66 The idea that group identification may engender exercise participation and adherence
67 is not new in itself, with research evidence, including meta-analysis, supporting this [14-16].
68 In general, participants reporting to identify more strongly with a group for which exercise
69 is a norm also report increased intention to exercise [16] and groups have been shown as
70 beneficial in terms of exercise adherence [14]. More recently, studies have adopted the
71 social identity approach as a theoretical standpoint. Stevens et al. [7,17] demonstrated a
72 relationship between levels of group identification and exercise attendance rates (in group
73 exercise classes and parkrun). Whilst undoubtedly promising research in terms of the
74 applicability of the social identity approach to group exercise contexts, these studies adopted
75 attendance measures of frequency/volume (e.g., Stevens et al. [7] measured percentage of
76 total sessions attended in a single week). To our knowledge, no studies to date have
77 specifically examined the social identity approach regards exercise classes and adherence in
78 terms of *tenure*, i.e. how long a person has attended a given class. Maintenance presents a

79 key area of difficulty in encouraging health behaviour change, such as exercise participation
80 [18]. By examining tenure, we have an opportunity to better understand factors relating to
81 continued adherence and encourage long-term participation. This may not be fully
82 comprehended through focus on frequency/volume of attendance in a single week.

83 A further opportunity is presented with regards the conceptualisation of group
84 identification. Beauchamp [19] highlighted the influential role of identification within
85 exercise groups as a way to promote adherence and advocated a need for greater
86 understanding. Whilst recent studies examining group identification in exercise classes have
87 used a global measure of identification [7,20], research evidence across disciplines has
88 suggested that group identification may in fact be multi-dimensional [21,22]. Cameron's [23]
89 three-factor model of social identity developed in social psychology has been empirically
90 supported in sport settings (e.g., [21]). The model proposes two cognitive elements of
91 identity (ingroup ties and cognitive centrality) and an affective element (ingroup affect).
92 Although an alternate form of physical activity to exercise, in a sport context support has
93 been shown for the three factors relating to different outcomes, such as initiative, and
94 pro/anti-social behaviour [21,24]. It is therefore plausible to posit that each factor may relate
95 differently to behaviours within an exercise context, for example, the behaviour of 'adhering
96 to an exercise class', meaning investigation may add novel and important insight.

97 **The Social Identity Approach to Leadership**

98 Group exercise class instructors may also have a key role to play in regard to
99 attendance [25] often planning and advertising classes as well as providing instruction. The
100 application of the social identity approach to physical activity has garnered interest in how
101 *leaders* (e.g. class instructors) in these settings may influence physical activity behaviour
102 [7,9,20,26]. Exercise leaders (i.e. instructors) may be an important influence on behaviour
103 of individuals who attend exercise classes (groups), and it has been suggested that the social

104 identity approach to leadership (SIL) could add substantial value to research in exercise
105 settings [6,7]. For instance, SIL places the group processes inherent within exercise class
106 settings at its analytical heart. According to SIL [27,28] effective leaders are: (1) in-group
107 prototypes ('prototypicality' - exemplifying the group and what makes it different); (2) in-
108 group champions ('advancement' - promoting the good of the group); (3) entrepreneurs of
109 identity ('entrepreneurship' - defining the shared beliefs and values); and (4) embedders of
110 identity ('impresarioship' - providing opportunities for the desired collective behaviour).
111 Through these four behaviours leaders are able to mobilise followers by recruiting and
112 maintaining the allegiance and support of other group members. In the context of exercise,
113 a class instructor may therefore be able to mobilise followers to attend classes. It is possible
114 to hypothesise that individuals attending exercise classes who perceive their instructor to
115 demonstrate identity leadership behaviours may be more likely to have greater class
116 adherence.

117 Until recently, there was limited direct research as to the relevance of SIL in exercise
118 settings. A small number of studies have now begun to examine the relationship between
119 both identity leadership behaviours and group identification in exercise participants [7,9,20].
120 Stevens et al. [7] found a small significant association between identity leadership
121 behaviours and attendance in exercise class attendees. However as previously noted, the
122 study measured attendance via proportional frequency as opposed to by tenure/adherence,
123 making predictions about applicability for long-term participation potentially difficult. In a
124 follow-up study, Stevens et al. [9] used a 2-wave design to show that perception of identity
125 leadership behaviours at study outset predicted group identification 8-weeks later and, that
126 this group identification was associated with attendance. These findings provided support
127 for the suggestion that identity leadership behaviours influence attendance via group
128 identification temporally. Yet, this study examined members of amateur sports teams - not

129 exercise classes. Support for SIL specifically within a group exercise context has recently
130 been offered by Steffens et al. [20], who demonstrated that identity leadership behaviours
131 were associated with more frequent class attendance (in the previous 4 weeks) via stronger
132 group identification. Thus, whilst evidence is growing that identity leadership behaviours
133 may engender exercise class attendance, and that this may occur via group identification,
134 empirical evidence regarding group exercise classes specifically is limited and offers scope
135 for further investigation, particularly in terms of adherence as measured by tenure, and
136 future intentions to continue to adhere.

137 **Relational Identification**

138 A distinct yet related concept to group identification is that of relational
139 identification [29,30]. Sluss and Ashforth [29] suggest that even within a group, we form
140 one-to-one role-relationships with others, and that the impact of these relationships has been
141 overlooked in leadership research. Sluss and Ashforth outlined relational identification as
142 the extent to which we define ourselves in terms of a given role-relationship. In the context
143 of an exercise class and the present study, an individual may identify themselves in the
144 context of the relationship between themselves, as the class attendee, and the class instructor
145 (i.e. leader). It is suggested that as two people get to know each other they become
146 'personalised' over time, thus affecting their interactions (behaviours) within the role-
147 relationship. Sluss and Ashforth propose that relational identification may lead to
148 identification with groups that are shared by the individuals in the role-relationship via
149 generalisation. Thus, relational identification may lead to group identification.

150 Sluss et al. [31] offered support for the association between relational and group
151 identification, and found that this was especially the case when the leader was viewed as
152 'prototypical'. In contrast, Steffens et al. [30] found that identification with a group may
153 lead to relational identification with a leader, and that the group provides the framework for

154 the relationship. It is plausible therefore that interaction in both directions may take place.
155 Whilst neither of the aforementioned studies pertained to an exercise situation, relational
156 identification to leaders has been shown to be associated with follower mobilisation of effort
157 in an elite sport context [32] and thus it is possible, but not yet known, that this may also
158 apply to exercise.

159 Since prototypicality is proposed as an element of identity leadership behaviour, its
160 potential involvement makes way for the possibility that group identification, identity
161 leadership, and relational identification may *all* be associated with, or predict, exercise class
162 adherence. As 'personalisation' takes place over time and influences behaviour, and with
163 potential multi-directional interactions at hand, a group exercise class attendee may both
164 identify more strongly with a class or leader they have attended/known for longer and/or
165 attend for longer when they identify more strongly with the leader or class. If both: (1)
166 relational identification with an exercise class instructor may affect behaviour/effort, and
167 lead to group identification with the exercise class; and (2) group identification with the
168 class may affect behaviour/adherence (through norms/values), and lead to relational
169 identification with the class instructor; we propose that it is a step forward, and logical, to
170 examine all three factors (group identification, identity leadership behaviours and relational
171 identification) together in a single study. Presently, this has not been examined.

172 **The Present Research**

173 With group exercise classes a potentially important way of encouraging exercise
174 participation, and the social identity approach an increasing area of interest with regards to
175 physical activity adherence and leadership, the present study examined the associations
176 between group identification, identity leadership, relational identification, and exercise class
177 adherence. We also aimed to examine the potential relevance of multi-dimensional factors
178 of group identification. The study partially replicated that of Stevens et al. [7] but expanded

179 upon it in four main ways. First, through measurement of class adherence by tenure and
180 expected future participation as opposed to attendance by volume. Second, through the
181 inclusion of relational identification as a potential predictor. Third, by inclusion of a multi-
182 dimensional measure of group identification in addition to a global measure, and fourth,
183 through focussing solely on group *exercise* class attendees.

184 Based on previous research, it was expected that participants' group identification
185 with the exercise class (H1a), perceived identity leadership behaviours of the class instructor
186 (H1b), and relational identification with the class instructor (H1c) would be associated with
187 the length of time participants had attended the exercise class. We also predicted that each
188 of the independent variables (respectively H2a, 2b, and 2c) would predict participant
189 expected likelihood of future attendance. These formed our primary hypotheses. A
190 secondary hypothesis was that the multi-dimensional elements of group identification may
191 show differing relationships with length of class attendance (H3).

192 It was anticipated that the findings may both support and offer new insights
193 regarding the potential role of identification in exercise adherence. From a practical
194 perspective, our research, and others', may ultimately contribute to recommended design
195 and marketing of exercise classes, as well as the training provided to class instructors to
196 promote long-term adherence of attendees.

197 **Method**

198 **Participants and Design**

199 Participants were 120 U.K. based adult exercise class attendees (male: $n = 13$, female:
200 $n = 107$) aged between 18 and 71 years ($M = 40.11$, $SD = 11.27$). Participants were part of a
201 range of exercise class types (see Table 1), and attended classes between "once every 2
202 weeks" and "more than 3 times per week", with the majority ($n = 61$) attending once per
203 week (see Table 2). Participants who attended classes with more than one regular instructor

204 were excluded from the study. A cross-sectional survey was conducted online and was open
 205 for participant completion from March to October 2019. Use of a cross-sectional survey
 206 design aligned with that used by Stevens et al. [7] allowing for partial replication and
 207 expansion of their findings. Participants took an average of 10 minutes and 34 seconds to
 208 complete the survey, which aligned with the researchers' expected timeframe.

209 Table 1.

210 *Participants by Exercise Class type*

211

Class type	No. Participants
Aerial Aerobics	1
Barre	1
Body Attack	1
Body Balance	5
Body Combat	3
Body Conditioning	1
Body Pump	4
Bootcamp	1
Boxercise/Box Fit	5
Circuits	9
Crew	1
CrossFit	1
Dance/Dance Fitness	2
Drum Fit	1
General Workout	1
HIIT	23
Insanity	2
Jazzercise	1
Kettlebells	2
Kickboxing	1
Legs, Bums and Tums	1
Over 50s Ballet	1
Pilates	19
Running	1
Spinning	15
Triple Challenge	1

Yoga	10
Zumba	6

212

213 Table 2.

214 *Participants by Frequency of class attendance*

215

Attendance Frequency	No. Participants
More than 3 times per week	5
3 times per week	15
Twice per week	35
Once per week	61
Once every two weeks	4

216

217 Power tables in Clark-Carter [33], indicated that for an R^2 value of 0.148 (based on
 218 Stevens et al.'s [7] study examining identity leadership, group identification, and exercise
 219 class attendance), with 3 predictors (representing; group identification, identity leadership
 220 behaviour, and relational identification) a sample size of 60 is required to achieve power of
 221 .80.

222 **Procedures**

223 The study was granted ethical approval by an Institutional Ethics committee. A small
 224 pilot study ($n = 4$) was conducted in March 2019. The process of collecting the participant
 225 data was without issue and procedures remained as planned. Voluntary and snowball
 226 sampling were used for participant recruitment. A link to an online survey was posted on
 227 social media (Facebook, Twitter, LinkedIn) and the primary researcher's own website.
 228 Participants could also share the link to reach other potentially suitable and interested
 229 participants.

230 Individuals were invited to visit the link which took them to the participant
 231 information sheet. They were then given the option to continue to the survey and by doing
 232 so, provided informed consent for participation. This method of consent aligned with that by

233 Stevens et al. [7] (where consent was assumed through survey completion) and British
234 Psychological Society [34] guidelines for low risk studies. Participants completed
235 demographics questions (age, gender, location, frequency of class attendance) and
236 inclusion/exclusion criteria questions. Participants were asked the type of class which they
237 would be answering the questions in relation to (if they attended multiple classes, they were
238 requested to answer in relation to one only) and to confirm that the class had one regular class
239 instructor. Participants subsequently completed the questionnaires; length of attendance to
240 date, expected likelihood of future attendance, group identification, social identity leadership
241 behaviour and relational identification. The final page of the survey acted as a participant
242 debrief.

243 **Measures**

244 **Group identification.** To assess participants' identification with their exercise group,
245 and replicate Stevens et al. [7], a four-item measure of social identification recommended by
246 Postmes et al. [35] was used. Participants rated four statements on a 7-point Likert scale
247 from 1 (*fully disagree*) to 7 (*fully agree*). The statements were 'I feel committed to my
248 exercise group', 'I am glad to be a member of my exercise group', 'Being a member of my
249 exercise group is an important part of how I see myself' and 'I identify with my exercise
250 group'. Higher scores indicated higher levels of identification with the group. This measure
251 has demonstrated good internal consistency in previous studies [7,20] and likewise did so in
252 the present study (Cronbach's $\alpha = .87$).

253 **Multi-dimensional group identification.** A multi-dimensional measure of
254 identification enabled investigation of H3. The social identity for sport questionnaire (SIQS;
255 [21]) has been developed in-line with theory and tested to enable measurement of social
256 identification by three factors; ingroup ties, cognitive centrality, and ingroup affect. The
257 scale's 9-items were adapted to fit the context of the study with the word 'team' being

258 replaced with 'exercise class' (e.g. 'I feel strong ties to other members of this exercise class',
259 'In general, I'm glad to be a member of this exercise class'). Participants rated the items to
260 indicate how they felt about being part of the exercise class from 1 (*strongly disagree*) to 7
261 (*strongly agree*). Higher scores indicated higher levels of the element of social identification
262 being measured. Although in an exercise class as opposed to sport team context, the use of
263 the SIQS was considered as having the potential to add insight, and be the first application of
264 this scale in an exercise context. In view of this, a three-factor (i.e., three subscales of the
265 SIQS) confirmatory factor analyses (via R software 4.0.2) using chi-square (χ^2), the robust
266 CFI, the robust Tucker-Lewis index, the robust root mean square error of approximation with
267 a 90% confidence interval (RMSEA), and the standardized root mean square residual
268 (SRMR; [36]) confirmed acceptable fit ($\chi^2 = 57.885, p < .001$, Robust CFI = .960, Robust
269 TLI = .940, Robust RMSEA = .108, SRMR = .054). All factor loadings were above .76. The
270 SIQS has been shown as being reliable with good construct validity [21]. The present study
271 illustrated good internal consistency for each of the three sub-scales; ingroup ties, cognitive
272 centrality, and ingroup affect; with Cronbach's $\alpha = .92, .87$, and $.91$ respectively.

273 **Identity leadership.** The Identity Leadership Inventory (ILI: [37]) assessed
274 participants' perceived SIL behaviour of their class instructor. Comprising 15-items, this
275 scale was developed to measure identity leadership, as per the theory, both globally and by
276 the four-facets proposed by Haslam et al. [27]: prototypicality, advancement,
277 entrepreneurship and impresarioship. Participants were asked to rate the extent to which their
278 exercise class instructor engaged in the behaviours/activities described by each of the fifteen
279 statements (e.g. 'the class instructor is a model member of the class', 'the class instructor
280 creates a sense of cohesion within the class') on a 7-point Likert scale from 1 (*not at all*) to 7
281 (*completely*). Higher scores indicated a higher level of engagement in identity leadership.
282 The ILI has demonstrated good construct validity both during development [37] and across

283 20 countries [28]. Identity leadership research in exercise settings has used the ILI [7] and
284 shown good internal consistency. Cronbach's α in the present study showed excellent
285 internal consistency ($\alpha = .94$) for both the global measure and the four facets ($\alpha = .89, .90,$
286 $.91, .82$ respectively).

287 **Relational identification.** To measure participants relational identification with their
288 exercise class instructor, we adapted a 10-item measure from previous studies [38,39]. Ten
289 statements (e.g. 'When I talk about my class instructor, I usually say 'we' rather than 'him or
290 her'' and 'The values of my class instructor are consistent to my own') were rated on a 7-
291 point Likert scale from 1 (*strongly disagree*) to 7 (*strongly agree*). Higher scores indicated
292 greater levels of relational identification.

293 Whilst the measure selected was intended for a workplace context, limited research on
294 relational identification in exercise settings rendered scale selection challenging. Slater et al.
295 [32] used a three-item measure of relational identification in an elite sports context.
296 However, this measure was intended primarily as a manipulation check and has not been used
297 elsewhere in the literature. The more widely used 10-item measure was therefore selected
298 and adapted to reference the 'class instructor' in place of 'supervisor'. The measure has
299 shown good internal consistency in previous research [38,39] and also did so in the present
300 study (Cronbach's $\alpha = .91$). Given that this was the first use of this measure of relational
301 identification in an exercise setting, unidimensional confirmatory factor analyses were run,
302 which evidenced non-acceptable model fit ($\chi^2 = 146.842, p < .001, \text{Robust CFI} = .846,$
303 $\text{Robust TLI} = .802, \text{Robust RMSEA} = .163, \text{SRMR} = .078$). Despite unsatisfactory fit, all
304 factors loadings were above .40 [40]. Nevertheless, relational identification results should be
305 interpreted with caution.

306 **Length of exercise class attendance to date.** Participants indicated how long they
307 had been attending the exercise class in relation to which they were answering the survey. It

308 was felt that 'weeks' attendance would be too granular for longer term attendees to recall and
309 report accurately, and that 'years' would not capture shorter term attendees. Thus, in an
310 attempt to balance obtaining data that was as accurate and granular as possible whilst
311 ensuring clarity for participant completion, *months* were selected as the attendance units.
312 Participants selected from a drop-down list ranging from 1 – 60 months, with additional
313 options of 'less than 1 month' and 'more than 5 years' (thus a total of 62 options were
314 available). Selection of either of the two latter options elicited a free-format follow up box in
315 which participants indicated the length of time they had attended. Data provided in this latter
316 circumstance was subsequently converted to months (e.g., 2 weeks = 0.5 months) prior to
317 analysis to ensure consistency of units.

318 **Expected likelihood of future attendance.** Measurement of anticipated future
319 participation in a cross-sectional study poses some issues since intention may not accurately
320 predict behaviour [41]. Additionally, there has been debate in the literature regarding the use
321 of 'intention' versus 'expectation' [42-45]. Meta-analysis by Sheppard et al. [44] indicated
322 that expectation was more highly correlated to behaviour than intention. In research directly
323 relating to physical activity (taking a walk) Warshaw and Davis [45] showed expectation to
324 be a significantly better predictor of behaviour than intention and more recently expectation
325 has been shown as more highly correlated to other health behaviours (alcohol consumption &
326 sustained weight loss) [42]. As such, *expected likelihood of future attendance* was
327 considered the most appropriate measure.

328 Courneya and McCauley [43] defined expectation as the "likelihood of the behaviour
329 being performed whether or not an intention exists" (p.53). Adapted from research
330 conducted by Spink [46] our participants were asked '*How likely are you to continue
331 attending the exercise class for the foreseeable future?*'. Fixed-choice responses available

332 were: (1) Not likely at all (at or near 0% chance); (2) Not likely (25% chance or less); (3) So-
333 So (50% chance); (4) Likely (75% chance or better); or (5) Very Likely (at or near 100%).

334 **Data Analysis**

335 To assess H1a, 1b, and 1c, a forced entry multiple regression was used to establish
336 whether group identification, identity leadership and/or relational identification were
337 associated with *length of class attendance*. To establish whether the three predictor variables
338 were associated with *expected likelihood of future attendance* (H2a, 2b, and 2c), an ordinal
339 regression was planned. Inspection of the data revealed the majority of participants (97%, n
340 = 116) had selected only two of the potential five answer categories ('at or near 100%' likely
341 and '75% or better'). Faced with this unexpected scenario and with an ordinal regression no
342 longer viable, a decision was taken to instead examine participants who were 'at or near
343 100%' likely to continue (n = 90) vs those who were not (n = 30 - comprising: 75% chance or
344 better, n = 26; 50% chance, n = 3; 0% chance, n = 1) using a binary logistic regression.
345 Although not ideal, given the focus on adherence, a distinction between participants who
346 reported being 100% likely to continue and those who did not, was felt to have potential
347 practical relevance. Forced entry was selected due to the limited amount of previous research
348 on exercise, none containing all three of the potential predictors to be investigated, and since
349 we were interested in which of the three predictors, group identification, identity leadership
350 or relational identification, was most strongly associated with class tenure. To address H3,
351 correlational analysis was used to assess whether the three SIQS elements showed differing
352 relationships with *length of class attendance*.

353 **Results**

354 **Preliminary Analysis**

355 Scatter plots for group identification, identity leadership and relational identification
356 (predictor variables) against length of class attendance (continuous outcome variable) were

357 run to ensure the relationships were not curvilinear. Z scores identified outliers as those with
 358 values over 3.29 [47]. Three values for length of class attendance and one value for identity
 359 leadership met the criteria and, the researcher being satisfied that this data was from the
 360 population being studied, were winsorized. P-P plots to check for normality revealed
 361 potential skew, particularly for identity leadership and length of class attendance. This was
 362 potentially not a concern due to the sample size being over 30 [47]. ZResid vs ZPred and
 363 partial plots revealed potential funnelling and a possible violation of homoscedacity with the
 364 identity leadership variable. The assumption of independent errors was checked using the
 365 Durbin-Watson statistic which was close to 2 (2.06). Collinearity between predictors was not
 366 considered an issue since the correlation coefficient between the predictors were all below .8,
 367 with the highest being between identity leadership and relational identification ($r = .68$).
 368 Additionally, the VIF values for the predictors were all well below 10 (highest = 2.114) and
 369 tolerance values above .2. The average VIF was 1.98 which was close to 1. Inspection of the
 370 residuals revealed no cases had a Mahalanobis distance of > 15 (largest = 13.419), none had
 371 Cooks distance > 1 (max = .151) and DFBETA statistics were all < 1 . It was concluded that
 372 extreme cases were unlikely to provide undue influence. The assumption of linearity
 373 between the predictor variables and their logits (required for binary logistic regression) was
 374 met since no interactions were significant ($p > .05$). The means and standard deviations of
 375 the continuous variables are shown in Table 3.

376 Table 3.

377 *Means and standard deviations of continuous variables*

378

	<i>M</i>	<i>SD</i>
Length of Class Attendance (months)	24.14	24.76
Group Identification	5.05	1.48

Identity Leadership	5.70	1.09
Relational Identification	4.66	1.31

379

380 **Main Analysis**

381 **Multiple regression.** Multiple linear regression analysis to test whether group
 382 identification, identity leadership, or relational identification significantly predicted
 383 participants' length of class attendance showed a non-significant regression equation
 384 ($F(3,116) = .795, p = .499$) with a small effect size ($r = .142$) (See Table 4). The three
 385 predictors explained 2.0% of the variation in length of class attendance ($R^2 = .020$). This
 386 suggests that other factors are involved since 98.0% is unexplained by the model. None of
 387 the three predictors significantly predicted length of class attendance; group identification (β
 388 $= .004, p = .974$), identity leadership ($\beta = .163, p = .222$), relational identification ($\beta = -.037,$
 389 $p = .780$). In sum, H1a, 1b, and 1c were not supported.

390 Table 4.

391 *Linear model of predictors of length of class attendance with 95% confidence intervals*
 392 *reported in parenthesis.*

393

	<i>B</i>	<i>SE B</i>	β	<i>P</i>
Constant	5.951 (-18.223, 30.125)	12.205		.627
Group Identification	.066 (-3.971, 4.103)	2.038	.004	.974
Identity Leadership	3.709 (-2.276, 9.693)	3.022	.163	.222
Relational Identification	-.707 (-5.717, 4.302)	2.529	-.037	.780

394 *Note. $R^2 = .020, p = .499$*

395

396 **Binary logistic regression.** Binary logistic regression tested whether group
 397 identification, identity leadership or relational identification could significantly predict

408 whether participants were at or near 100% likely to continue attending their exercise class for
 409 the foreseeable future vs less than at or near 100% likely to continue (H2). Results showed a
 410 significant regression equation ($chi-square = 12.918, df = 3, p = .005$) indicating that the
 411 model provided a significantly better prediction than those made purely on which category
 412 occurred most frequently. The model explained 15.1% of the variation in outcome ($pseudo$
 413 $R^2 = .151$) and the Hosmer and Lemeshow test indicated a good fit ($chi-square = 6.531, df =$
 414 $8, p = .588$). Group identification showed a non-significant positive effect ($b = .301, Wald =$
 415 $2.443, df = 1, p = .118$), identity leadership showed a significant positive effect ($b = .632,$
 416 $Wald = 4.926, df = 1, p = .026$), and relational identification showed a non-significant
 417 negative effect ($b = -.324, Wald = 1.685, df = 1, p = .194$) on likelihood of continued
 418 participation being at or near 100% (see Table 5). H2b was therefore supported but H2a and
 419 2c were not. Table 5 summarises the results.

420 Table 5.

421 *Model of predictors of likelihood of future participation with 95% confidence intervals*
 422 *reported in parenthesis.*

423

	<i>B</i>	SE B	<i>P</i>	Odds Ratio
Constant	-2.397	1.133	.034	0.091
Group Identification	.301	.193	.118	1.352 (.926, 1.973)
Identity Leadership	.632	.285	.026	1.882 (1.077, 3.288)
Relational Identification	-.324	.249	.194	.724 (.444, 1.179)

424 *Note. Pseudo $R^2 = .151, p = .005$*

425

426 The analysis indicated that for every increase of 1 unit in identity leadership score it is
 427 1.9 times more likely that someone will be 'at or near 100%' likely to continue attending
 428 their exercise class for the foreseeable future versus not. The confidence interval for the odds

419 ratio did not span 1, indicating the direction of the observed relationship is likely to be true in
 420 the population sampled. The confidence intervals for the odds ratios of group identification
 421 and relational identification did span 1, providing ambiguous directional results.

422 **Correlation.** Pearson's correlation coefficients were used to examine whether there
 423 were differences in the associations between the three elements of multi-dimensional group
 424 identification measured by the SIQS and the length of class attendance to date (H3). Results
 425 showed a small effect approaching significance for ingroup ties ($r = .165, p = .072$), a
 426 negligible non-significant negative effect for cognitive centrality ($r = -.051, p = .579$) and a
 427 negligible non-significant effect for ingroup affect ($r = .010, p = .910$). This provided partial
 428 support for H3 as there appeared to be marginal differences, both in effect size and direction,
 429 across the elements.

430 Table 6.

431 *Correlation matrix of study variables*

Variable	1	2	3	4	5	6	7	8
1. Group identification	-	.66**	.70**	.73**	.60**	.61**	.08	.30*
2. Ingroup ties		-	.51**	.57**	.63**	.48**	.17	.19*
3. Cognitive centrality			-	.63**	.47**	.47**	-.05	.19*
4. Ingroup affect				-	.56**	.54**	.01	.35**
5. Identity leadership					-	.68**	.14	.32**
6. Relational identification						-	.08	.16
7. Exercise to date (months)							-	.26*
8. Future likelihood								-

432 *Note.* * $p < .05$; ** $p < 0.001$.

433 *Future likelihood coded as 0 = Not likely at all (at or near 0%); 1 = So-so (50% chance); 2*
 434 *= Likely (75% chance or better); and 3 = Very likely (at or near 100%).*

435 **Exploratory Analysis**

436 **Binary logistic regression.** Given that elements of multi-dimensional group
 437 identification as measured by SIQS showed potential to have differing relationships with
 438 length of class attendance to date, and that identity leadership appeared to be associated
 439 expected likelihood of future attendance, a binary logistic regression was conducted to
 440 establish if the three elements of SIQS might significantly predict likelihood of future
 441 attendance. The assumption of linearity of the logits of the predictors was met, as was the
 442 assumption of multicollinearity (Tolerance values all $> .1$, $VIF < 10$, $r < .8$). P-P plots
 443 appeared to indicate normality, with possible skew for ingroup affect. The regression showed
 444 a significant equation ($chi-square = 12.397$, $df = 3$, $p < .01$) indicating that the model
 445 provided a significantly better prediction than those made purely on which category occurred
 446 most frequently. The model explained 14.5% of the variation in outcome ($pseudo R^2 = .145$)
 447 and the Hosmer and Lemeshow test indicated a good fit ($chi-square = 9.167$, $df = 8$, $p =$
 448 $.328$). Ingroup affect showed a significant and positive effect ($b = .672$, $Wald = 7.500$, $df =$
 449 1 , $p = .006$), ingroup ties showed a non-significant negative effect ($b = -.079$, $Wald = .174$, $df =$
 450 1 , $p = .677$), and cognitive centrality showed a non-significant negative effect ($b = -.059$,
 451 $Wald = .1081$, $df = 1$, $p = .742$) on likelihood of continued participation being at or near
 452 100% (see Table 7).

453 Table 7.

454 *Model of predictors of likelihood of future participation with 95% confidence intervals*
 455 *reported in parenthesis.*

456

	<i>B</i>	SE B	<i>P</i>	Odds Ratio
Constant	-2.011	.995	.043	0.134
Ingroup Ties	-.079	.188	.677	0.924 (.639, 1.337)
Cognitive Centrality	-.059	.179	.742	.943 (.664, 1.339)

Ingroup Affect	.672	.245	.006	1.957 (1.210, 3.165)
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457 *Note. Pseudo R² = .145, p < .01*

458

459 The analysis indicated that for every increase of 1 unit in ingroup affect score it is 2.0
 460 times more likely that someone will be 'at or near 100%' likely to continue attending their
 461 exercise class for the foreseeable future versus not. The confidence interval for the odds ratio
 462 did not span 1, indicating the direction of the observed relationship is likely to be true in the
 463 population sampled. The confidence intervals for the odds ratios of ingroup ties and
 464 cognitive centrality did span 1, providing ambiguous directional results.

465

Discussion

466 The goal of our study was to provide further support for the applicability of the social
 467 identity approach to physical activity, specifically in the context of group exercise class
 468 adherence and leadership. Additionally, we expanded on previous research [7,20] through
 469 examination of the potential relevance of relational identification and multi-dimensional
 470 factors of group identification. We sought to discover whether group identification, identity
 471 leadership and/or relational identification may predict exercise class adherence, as measured
 472 by length of class attendance to date and/or expected likelihood of future participation. We
 473 found that neither group identification, identity leadership behaviour nor relational
 474 identification were associated with length of class attendance to date. H1a, 1b and 1c were
 475 not supported. We did find support for H2b, as identity leadership behaviour was positively
 476 associated whether exercisers reported being 'at or near 100%' likely to continue attending
 477 their class vs not. Since neither group nor relational identification were associated with
 478 continuation likelihood, H2a and 2c were not supported.

479 The results appear to be at odds with previous work [7,9,20] which has suggested that
 480 identity leadership behaviour may influence group physical activity attendance via group

481 identification. The fact that neither of these variables were shown to be associated with the
482 length of exercise class attendance to date in the present study may shed a slightly different
483 light. Whilst mediational analysis was not conducted in the present study, this alone would
484 not likely account for the lack of findings. One explanation of our findings could relate to
485 how we examined attendance. Previous studies have examined attendance by
486 volume/frequency as opposed to *tenure*. Our findings open the possibility that *length* of
487 exercise class attendance and *volume* of attendance when a person is minded to attend may be
488 distinct exercise adherence related constructs, each with differing relationships to group
489 identification, identity leadership, and/or relational identification. Equally, it could be that
490 these predictors are more impactful or relevant in a sport context compared to exercise or that
491 exercise class leaders are less aware of behaviour that may engender group identification, due
492 to the lack of training and research in the area. Future longitudinal studies involving the
493 development and application of interventions (e.g., the 3Rs: Reflect, Represent, Realise; [27])
494 aimed at increasing both identity leadership behaviour and group identification, whilst also
495 measuring group identification and attendance both via frequency *and* length as well as
496 examining drop-out rates may help to provide a clearer understanding.

497 Identity leadership behaviour was however shown to be a significant predictor of
498 participants' expectations being 'at or near 100%' likely to continue attending their class.
499 For every unit increase in identity leadership score participants were 1.9 times more likely to
500 be 'at or near 100%' likely to continue vs not. This finding partly aligns with the previous
501 research evidence that has shown identity leadership behaviour to be associated with more
502 frequent attendance in both exercise classes [20] and amateur sports teams [9]. However, it
503 does not align with other results in the current study that identity leadership did not predict
504 length of class attendance. 'Expectation' of future behaviour has been shown correlate more
505 closely to actual behaviour in a variety of health behaviours [42,45] and as such the finding

506 that identity leadership was associated with greater expected likelihood of future
507 participation, specifically in exercise class attendees, was as predicted, and accordingly, has
508 practical relevance (although it is important to remember this particular study does not permit
509 causal establishment). Since group exercise classes offer a gateway to exercise [4] the
510 possibility that identity leadership behaviour might enhance expected or actual continuation
511 of participation could have important practical implications for increasing physical activity
512 adherence. Further research in this regard appears warranted as it may offer an important
513 angle for inclusion of identity leadership behaviours in exercise instructor training.

514 The finding that relational identification with the exercise instructor did not appear to
515 be related to adherence in the present study was not as expected. Relational identification
516 was not found to be associated with either length of class attendance to date or future
517 expected participation. Although this could indicate that relational identification with the
518 class instructor is not a relevant factor in the adherence of exercise class attendees, it could
519 equally be that the measure used was not suitable for an exercise context. Measure selection
520 was limited and that used had only been done so previously in workplace settings [38,39].
521 The development and validation of a relational identification measure specifically for
522 exercise class settings may help to enhance understanding and future work in this area.
523 Given the strong theoretical links to group identification and identity leadership, the null
524 finding in this study alone would seem insufficient to rule out a potential role for relational
525 identification regards exercise adherence.

526 Whilst no significant results were shown in relation to the global measure of group
527 identification, the study additionally aimed to explore the possibility that multi-dimensional
528 elements, as measured by the SIQS [21] may relate differentially to exercise class adherence.
529 It appeared that this may be the case. Initial correlational analysis showed that the three
530 elements had differing relationships with length of class attendance and whilst none were

531 significant, small effects were observed for ingroup ties. Of particular interest was follow-up
532 analysis indicating that ingroup affect (which specifically measures how good people feel
533 about being a member of a group) was significantly associated with whether participants
534 were 'at or near 100%' likely to continue vs not. For every unit increase in ingroup affect
535 score participants were 2.0 times more likely to be 'at or near 100%' likely to continue to
536 exercise in the class. This finding, whilst exploratory in nature, expands on previous research
537 indicating the importance of group identification on exercise class attendance by highlighting
538 that it may be the specific element of identity relating to how good a person feels about being
539 a member of a group that could enhance expected adherence. This aligns with research in the
540 field of sport suggesting that the different elements measured by SIQS may relate to different
541 outcomes [21,24]. Clearly the small scale and cross-sectional nature of this study does not
542 enable any firm conclusions, but does suggest that further research on exercise adherence
543 utilising SIQS might enhance present understanding.

544 **Further Limitations and Considerations**

545 Whilst the present study has yielded some interesting and perhaps unexpected results,
546 there are limitations to note. First, the sample examined may not have been representative of
547 the group exercise class attendee population and may not be generalisable due to voluntary
548 and snowball sampling. Indeed, it is possible that only those more enthusiastic about their
549 class responded. This was borne out by the majority of participants (97%) selecting ratings
550 of either 'at or near 100%' or '75% or more' likely to continue. An alternative sampling
551 approach whereby classes are approached as a whole and attended in person by the
552 researcher, such that all attendees from a class are given an opportunity to participate, may
553 help address this were the study to be repeated. This is partially akin to the approach adopted
554 by Stevens et al. [7] and could explain the difference in findings. A systematic approach with

555 sufficient sample sizes could allow researchers to examine differences across exercise class
556 types.

557 Further regarding generalisability, the participants comprised primarily females (89%)
558 affording the possibility that the findings may not be as applicable for male class attendees.
559 The greater number of female respondents was perhaps unsurprising, since research has
560 suggested that female:male class attendance is in the region of 5:1 [48]. From a practical
561 standpoint, since more males than females are presently meeting physical activity guidelines
562 [49] opportunities to promote exercise activity by females may be of particular importance.

563 Third, our study was cross-sectional and relied on self-report data. We asked
564 participants how long they had attended the exercise class, but how objectively accurate this
565 was is unknown. The use of a longitudinal design may partially eliminate this issue since
566 attendance could be recorded and measured over a known period. Coupling this with
567 attendance register records, may also add rigor. The cross-sectional nature of this research
568 means that whilst it may be speculated that group identification by way of ingroup affect may
569 positively affect group exercise class adherence (expectation to continue), no causal link can
570 be established. Instead, our research should be seen as an important stepping stone in terms
571 of examining the multi-dimensional conceptualisation of group identification in exercise
572 settings.

573 Fourth, in an attempt to partially replicate and expand on Stevens et al. [7], we elected
574 to use length of exercise class attendance to date as a dependent variable, in place of exercise
575 class attendance frequency. The aim of this was to allow comparison between results from
576 volume versus length of attendance data in a comparable cross-sectional study. In view of
577 this we chose not to additionally use length of class attendance as a control variable in the
578 first step of the expected future attendance regression. Controlling for attendance length in a
579 future study of expected class attendance continuation may provide additional insight.

580 Finally, although SIL behaviour showed a significant result in terms of predicting
581 likelihood of continued participation, mis-matched group sizes meant that 75% would have
582 been predicted correctly through chance alone. Binary logistic regression was undertaken in
583 place of the planned ordinal regression since most participants selected only two of the five
584 available likelihood categories, with greater number choosing 'at or near 100%'. Thus,
585 whilst the regression model with predictors offered a significant improvement, the prediction
586 percentage increased by just 2% to 77%. For ingroup affect, this was even smaller, from
587 75.0% to 75.8%. Further research, ideally with more equal group numbers, would need to be
588 conducted to check for replication.

589 **Conclusions**

590 The present study has provided partial support for a proposed role of SIL behaviour in
591 the continued participation of exercise class attendees. Additionally, we present evidence
592 that supports the proposition that the multi-dimensional elements of group identification may
593 be associated differently with exercise class adherence. Ingroup affect appears most pertinent
594 in this regard. Our study has also raised questions over whether the social identity approach is
595 as applicable for exercise class adherence in terms of *tenure* as has previously been suggested
596 in relation to attendance *frequency/volume* and indicates further investigation is warranted.
597 The lack of support for a role of relational identification regards exercise class adherence in
598 the present research should not preclude its investigation in future. Increasing activity levels
599 amongst adults is of paramount importance, and we have provided initial evidence that the
600 perceived leadership of 'our' exercise instructors, and ingroup affect may be determinants of
601 whether 'we' will continue to attend our exercise classes.

602

603 The Authors declare that there is no conflict of interest.

604

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